

# Examples of Instrumental Variable Analyses

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# Outline

- Finding an instrumental variable
- IVs in randomized trials: Moving To Opportunity
- Answering a parallel question with a natural experiment (lottery)
- IVs from natural experiments: compulsory schooling law changes
- IVs from genes: FTO as an IV for maternal obesity
- Goals:
  1. Recognize contexts in which IV analyses might be feasible and useful
  2. Recognize the limitations and assumptions of the IV analysis

# How do you find an Instrumental Variable?

1. Randomize
2. Some other possible sources of exogenous variation:
  - a. Geography of city
  - b. Policy variations
  - c. Institutional features (banking policies, loan guarantees)
  - d. Timing of newly available resources
  - e. Wait lists or lotteries for subsidies
  - f. Genetic polymorphisms

Randomizing is generally preferable, because the IV assumptions are more plausible and the 1<sup>st</sup> stage effects are often larger.

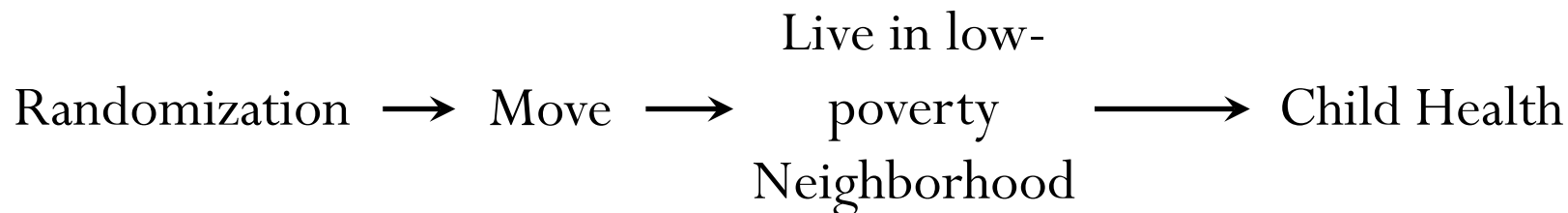
# Example 1, Moving To Opportunity Trial

- Families with children in urban public housing developments invited and randomized to:
  - Control
  - section 8, or
  - “low poverty” section 8 (must move to neighborhood with <10% poverty)
- Once randomized:
  - 60% of section 8 group moved
  - 47% of low poverty group moved.
- This may sound bad, but compare to a drug-based trials:
  - Women’s Health Initiative: “At the time the trial was stopped, 54.0% of study participants assigned to receive CEE and 53.5% of those assigned to receive placebo had discontinued use of their study medication.” –Hsia 2006
  - TODAY: “Adherence to the medication regimen before the primary outcome was reached or the study was completed ranged from 84% at month 8 to 57% at month 60” -TODAY study group, NEJM 2012

# Example 1, Moving To Opportunity

Multiple causal questions one might try to address with data from the Moving To Opportunity (MTO) trial:

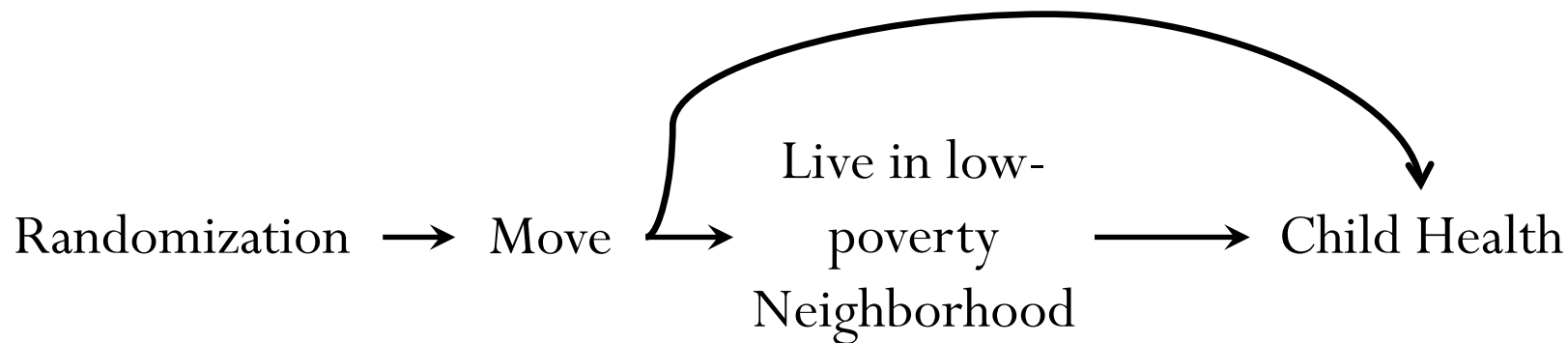
1. Does moving from very high poverty public housing developments benefit the health of mothers or their children?
2. Does living in a low poverty neighborhood benefit the health of mothers or their children?



# Example 1, Moving To Opportunity

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# Design

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This is the “first stage” estimate if you think of moving from the development as the endogenous variable.

# Did the trial affect neighborhood environment?

Poverty Rate	Control	ITT (Low Poverty)	
	Mean	Difference	P-value
Baseline	53.1%	-0.4	0.41
At 1 Year	50.0%	-17.1	<.001
At 5 Years	39.9%	-9.9	<.001
At 10 Years	33.0%	-4.9	<.0001

This is the “first stage” estimate if you think of neighborhood poverty as the endogenous variable.



# IV analyses in MTO

- Standard 2-stage least squares
- In most IV analyses, we think the “treated” group includes some “always treated” people and some “compliers”.
- The IV estimate refers to effect in the “complier” subgroup who received treatment because of the value of the IV.
- However, primary analyses of MTO define the endogenous variable as moving from the development *with the voucher given by the trial*.
- In this definition of the treatment, it is impossible to be treated if you are not randomized to receive a voucher.
- Therefore, everyone who is “treated” is a “complier” and the IV effect estimate = effect of treatment on the treated (TOT)

# Whose Causal Effect?

Response if assigned to receive a voucher:

Response if  
assigned to  
not receive a  
voucher:

	Don't Move	Move
Don't Move	Never-Takers	Compliers
Move	Contrarians/ Defiers	Always Takers

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# Early Results for Behavioral Problems, Boston 2 Year Low Poverty Group vs Controls

	<b>Control Mean</b>	<b>ITT Difference (SE)</b>	<b>TOT/IV Difference (SE)</b>
Boys	.326	-.090 (.041)	-.184 (.088)
Girls	.193	-.023 (.030)	-.046 (.056)

# Mid-Term (5-7 year) Results for Children's Mental Health (K6)

	<b>Control Mean</b>	<b>ITT Difference (SE)</b>	<b>TOT/IV Difference (SE)</b>
Boys	-.162	.069 (.091)	.167 (.223)
Girls	.268	-.246 (.091)	-.508 (.060)

# Trial challenges

- Mixed effects, attributable to
  - Small samples?
  - Heterogeneous effects?
- Uncertainty about the salient component of the treatment
  - Social disruption associated with moving?
  - Changes in residential environment?
  - Changes in schooling?
- Who are the compliers?

Most of these issues arise whether you use IV or ITT to analyze the data

# Example 1a:

- Causal question:

Does moving from very high poverty public housing developments benefit the health of mothers or their children?

We did a trial, but do you believe the results?

Can we get more evidence?

Voucher lottery

# Jacob & Ludwig 2011

- We match mortality data to information on every child in public housing that applied for a housing voucher in Chicago in 1997(N=11,848).
- Families were randomly assigned to the voucher wait list, and only some families were offered vouchers.
- Families randomized to the voucher moved to census tracts with an average of 7 points lower poverty.



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# Jacob & Ludwig 2011

- Treatment group = children whose families were assigned a waitlist number from 1 to 18,110, and so were offered a voucher by May 2003
- Control group = everyone assigned a higher lottery number.
- OLS with a person-quarter panel dataset for 1997:Q3 through 2005:Q4
- $y_{it}$  measures child  $i$ 's outcome in quarter  $t$ ,  $PostOffer_{it} = 1$  if child  $i$ 's family was offered a voucher prior to  $t$ , else  $PostOffer_{it} = 0$
- $X$  = control variables (whether the family is offered a voucher some time *after* quarter  $t$ , gender, splines for baseline age (kinks at 1, 2, 5, 8 and 15) and calendar time (kinks every 6 calendar quarters). Clustered standard errors.

# Jacob & Ludwig 2011

- ITT:

$$y_{it} = \alpha + \beta_1(PostOffer_{it}) + \mathbf{X}\Gamma + \varepsilon_{it}$$

- IV:

$$Leased_{it} = \alpha + \theta_1 PostOffer_{it} + \mathbf{X}\Gamma + \gamma_t + \varepsilon_{it}$$

$$y_{it} = \alpha + \pi_1 \widehat{Leased}_{it} + \mathbf{X}\Gamma + \gamma_t + \varepsilon_{it},$$

# IV analyses of a housing voucher lottery

		ITT	IV
<b>Boys Only</b>			
Death All Causes	100.32	60.44 (-51.52, 172.4)	203.88 (-175.12, 582.92)
<b>Girls Only</b>			
Death All Causes	38.16	-40.88 (-65.28, -16.52)	-130.20 (-207.6, -50.4)

Same analytic approach to natural experiment generated by a lottery and randomized experiment. Similar message re gender effect modification. Note large CIs.

# Example 2, natural experiment based on policy change

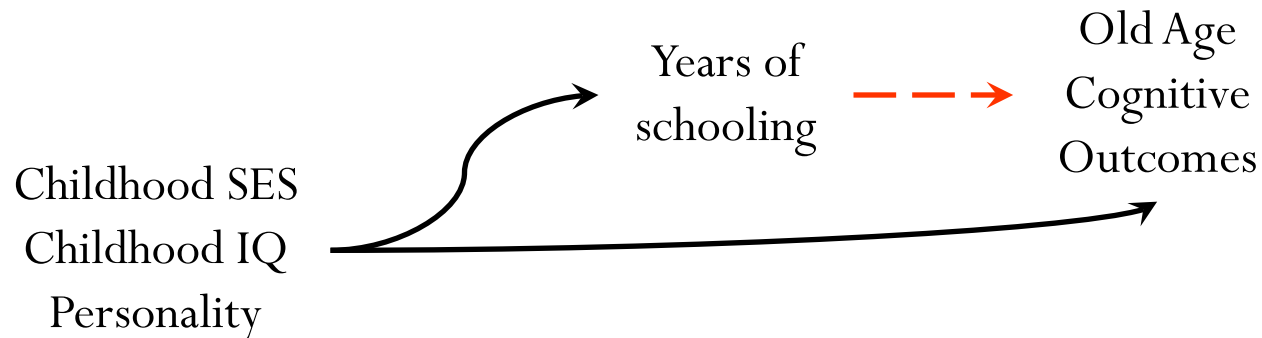
- Causal question:

Does completing additional years of education improve memory in old age?

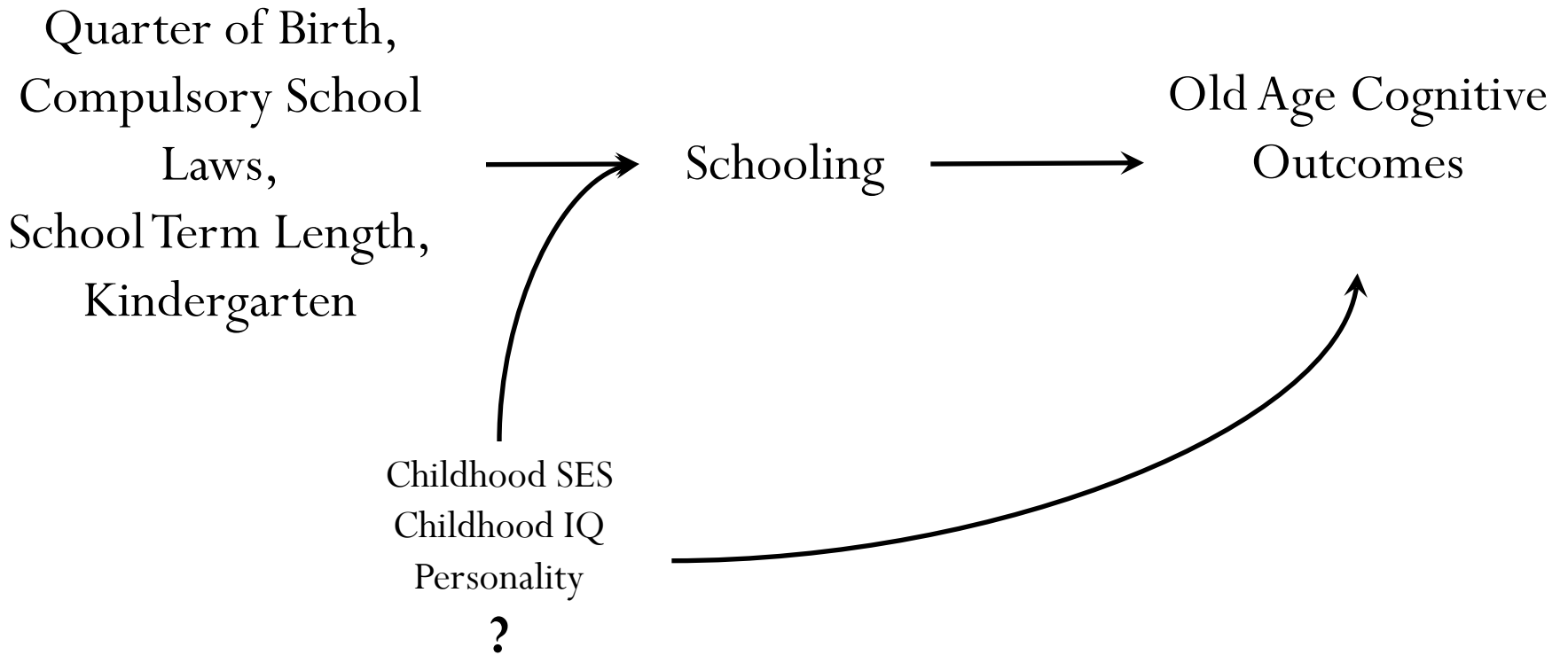
# Substantive Question

Multiple studies show that years of education predicts old age cognitive function, cognitive change, and dementia.

Causality questionable.



# Natural Experiments for Education





# Natural Experiments: UK Education Reform Effect on Education

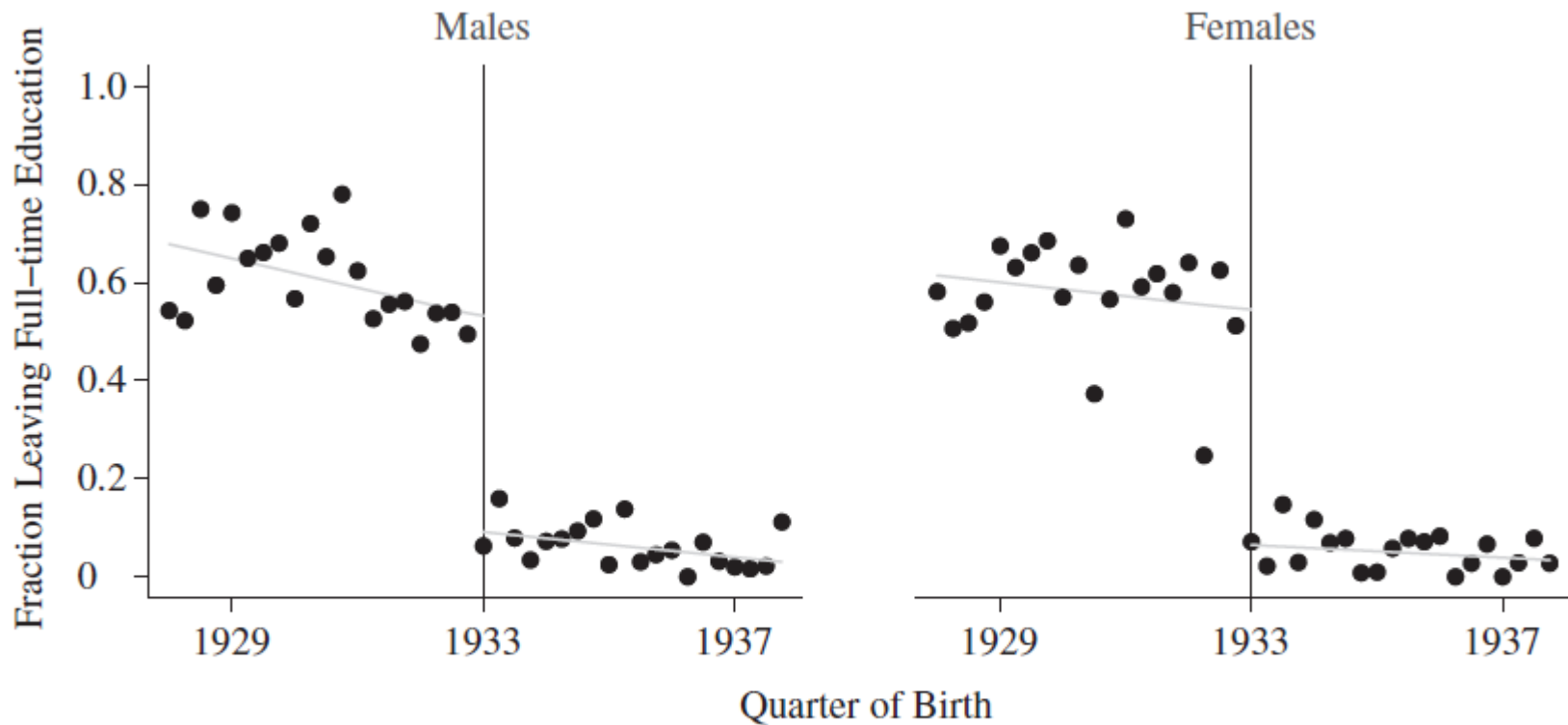


Fig. 1. *Effect of 1947 Reform on Fraction Leaving Full-time Education at or Before Age 14*

# Natural Experiments:

## UK Education Reform Effect on Education

**Reform had a powerful and immediate effect on about half the population of 14 years olds.**

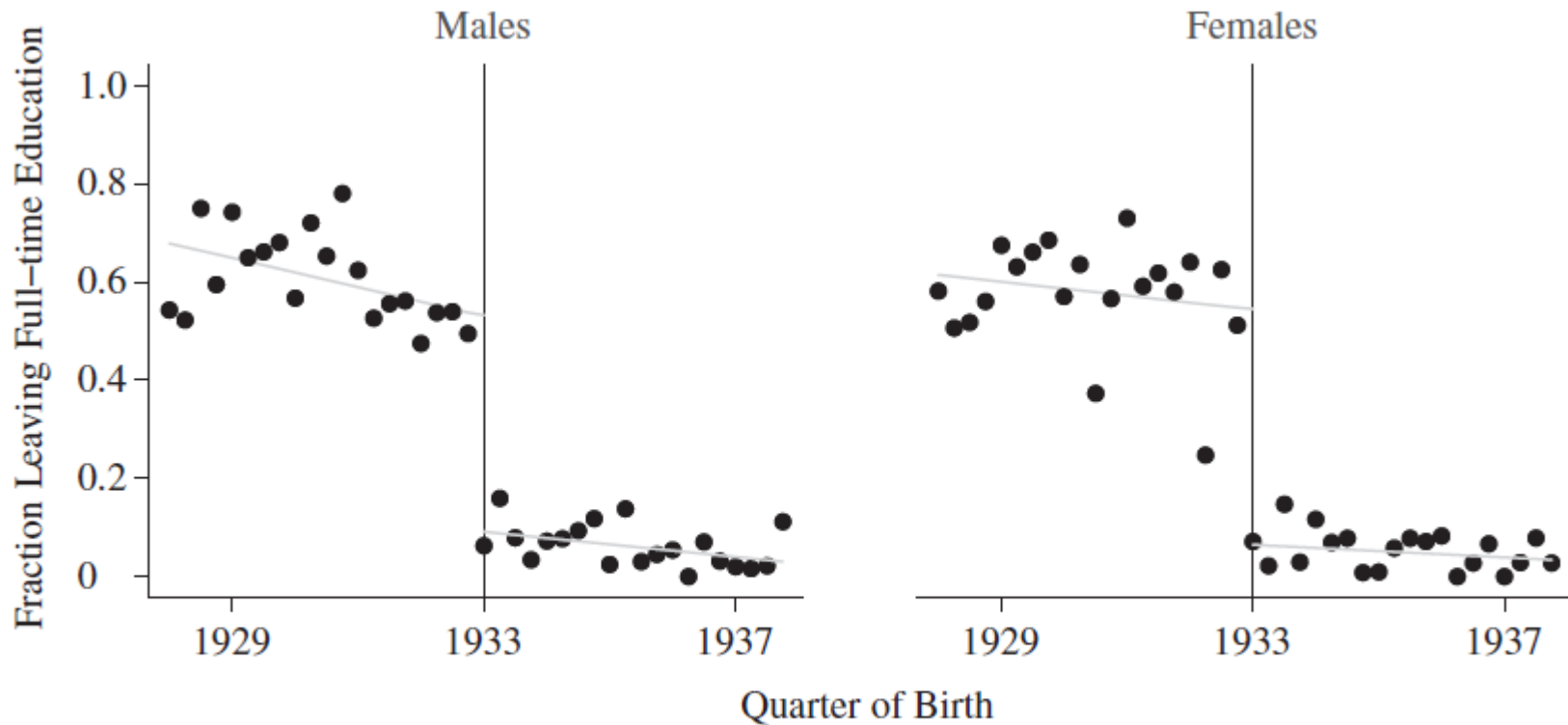


Fig. 1. *Effect of 1947 Reform on Fraction Leaving Full-time Education at or Before Age 14*

# Natural Experiments: IV Estimates for Education effect on EF

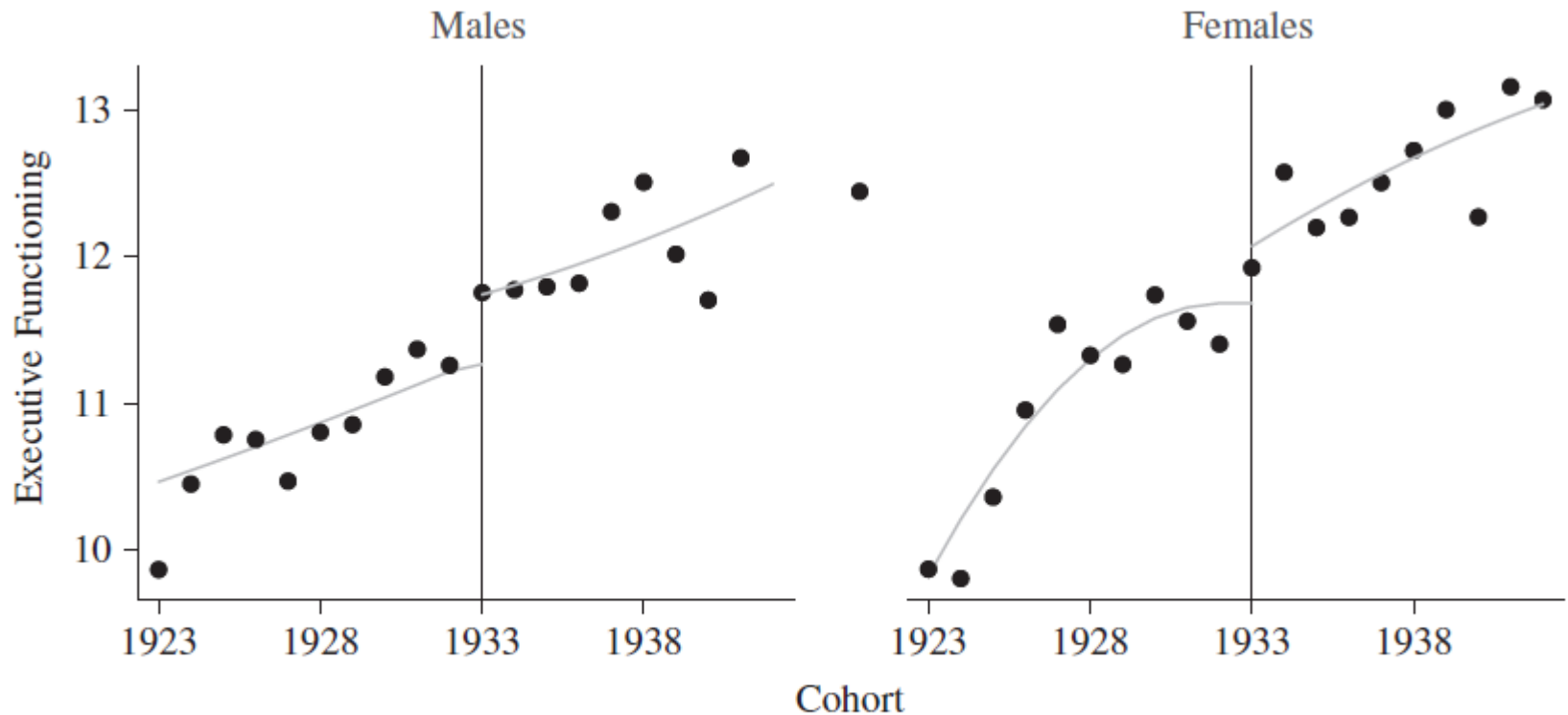


Fig. 9. *Effect of 1947 Reform on Executive Functioning (Conditional on Leaving Before 16)*

# Natural Experiments: IV Estimates for Education Effect on EF

**Note sensitivity to model for temporal trends.**

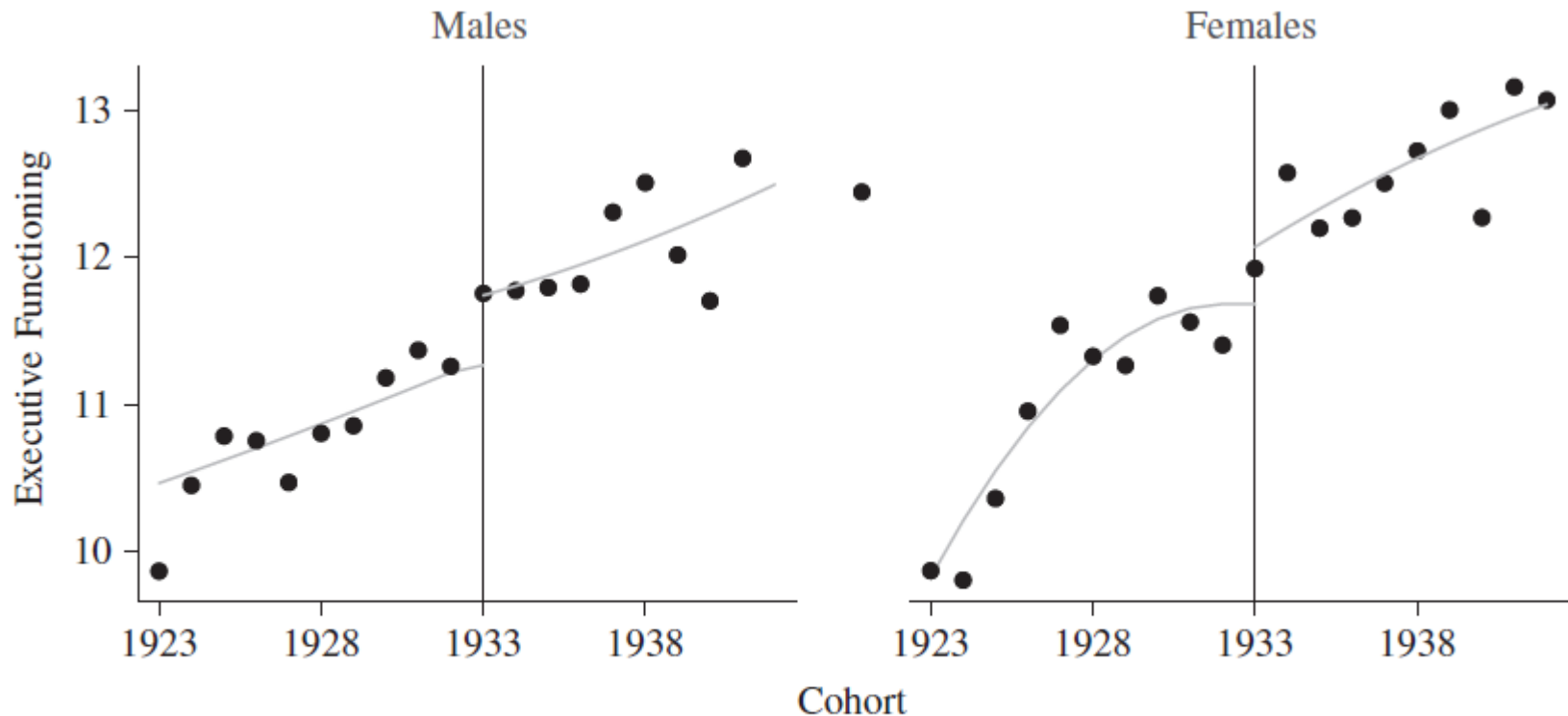


Fig. 9. *Effect of 1947 Reform on Executive Functioning (Conditional on Leaving Before 16)*

# Estimating the IV effect

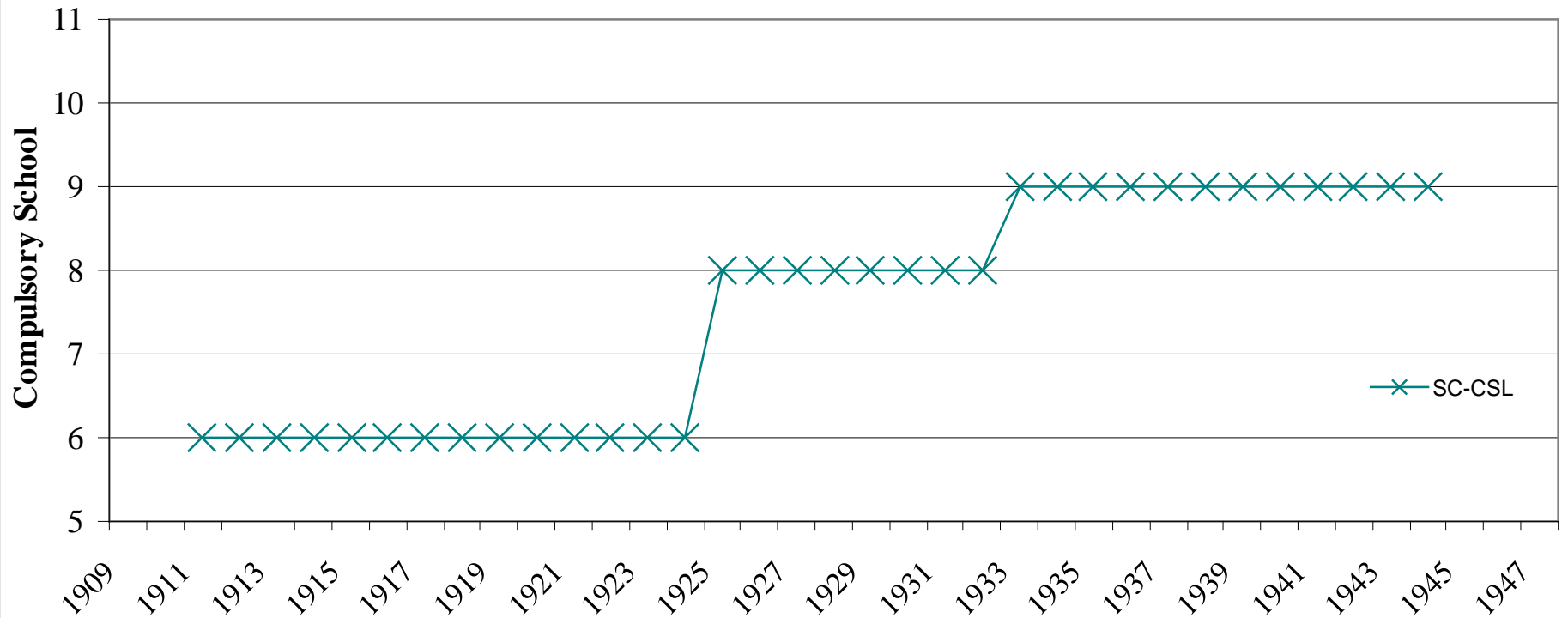
- Banks & Mazzone call this a “fuzzy regression discontinuity design” and estimate with 2SLS.

	Males			Females	
	Year band=1	Year band=3		Year band=1	Year band=3
Memory	.60 (.35)	.43 (.19)		.51 (.34)	.35 (.19)
Exec Fx	.64 (.36)	.37 (.19)		-.10 (.39)	.09 (.21)

# IV Estimates Using US Policy Changes

- Banks and Mazzona replicated earlier findings in the US
- Advantage of the US context:
  - Education is decentralized, so there were more places that changed policies
  - Allows for better control of secular trends: you can rule out a sudden change in 1947.
- Disadvantage of the US context:
  - Effect of the laws was very small
  - Generally not well enforced, most people would have attended more school than required anyway
  - Complier group is small.

# Early 20<sup>th</sup> Century CSL Changes



# IV Analyses

- State schooling policies
  - Compulsory school to drop out (CSL) or receive a work-permit (CSL-W)
  - Based on policy in state of birth when school-age
  - 2-Sample least squares analysis
- Exposure (endogenous) variable:
  - Years of education (self-report)



# Data Set: 1<sup>st</sup> Stage

- IPUMS (Census) 5% 1980 sample,
- Birth years 1900-1947
- Years of education linked to CSLs and CSL-Ws based on state of birth
- Link predictions from 1<sup>st</sup> stage regression model to individual data in the 2<sup>nd</sup> stage based on state of birth and all covariates.

# Data Set: 2<sup>nd</sup> stage

- Health & Retirement Study, 1992-2000: panel enrollment by birth cohort (whites only due to evidence on enforcement)
- Cognitive assessments and state of birth on 21,041 individuals born 1900-1947
- CSLs and CSL-Ws

# Two-Sample Least Squares

**Sample 1:**  
5% Census  
sample.

CSLs in  
each state  
and year,  
1906-1961.

**Stage 1:**  
Regress  
education  
on CSLs,  
with other  
covariates.

Predicted  
education  
( $\hat{E}$ ).

**Sample 2:**  
HRS data.

**Stage 2:**  
Regress health  
outcomes on  $\hat{E}$ ,  
with other stage 1  
covariates.  
Regression  
coefficient for  $\hat{E}$   
is the IV effect  
estimate.

# Covariates

- 1
  - Unadjusted
- 2
  - Sex
  - Birthyear (indicators for every year)
- 3
  - State of birth indicators
- 4
  - State characteristics: age 6 % black, % urban, and % foreign born; age 14 manufacturing jobs per capita and wages per manufacturing job

# Do the Instruments Predict Education?

## First stage regression results (from IPUMS 5% sample)

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	<b>1. Unadjusted Model</b>	<b>2. Birthyear and sex</b>	<b>3. Model 2 + state of birth</b>	<b>4. Model 3 + state condns</b>
CSLs	0.238 (0.236, 0.240)	0.110 (0.108, 0.112)	0.062 (0.059, 0.064)	0.037 (0.034, 0.040)
CSL-Ws	0.143 (0.146, 0.141)	-0.032 (-0.034, -0.029)	0.063 (0.060, 0.066)	0.044 (0.040, 0.048)
CSL-Ws UNR	-1.397 (-1.429, -1.365)	-0.282 (-0.315, -0.249)	-0.204 (-0.238, -0.17)	0.034 (0.000, 0.069)

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# How Strong is the 1<sup>st</sup> Stage?

	1. Unadjusted Model		2. Birthyear* and sex.		3. Model 2 + state of birth indicators		4. Model 3 + state characteristics <sup>#</sup>	
	$\beta$	95% CI	$\beta$	95% CI	$\beta$	95% CI	$\beta$	95% CI
Model $r^2$ without instrumental variables		0.0000		0.1080		0.1599		0.1626
Model $r^2$ including instrumental variables		0.0465		0.1127		0.1613		0.1631
Variance explained by instrumental variables		0.0465		0.0047		0.0014		0.0005

Not technically “weak” instruments, but clear that a small violation of the IV assumptions could introduce a large amount of bias.

# IV Estimates for Education: CSLs

Estimated effect of 1 year ed'n on cognitive test scores.

<b>Model covariates</b>	<b>Memory</b>		<b>Cognition</b>	
	$\beta_{IV}$	95% CI <sup>^</sup>	$\beta_{IV}$	95% CI <sup>^</sup>
1. Unadjusted	0.33	(0.27, 0.39)	0.19	(0.12, 0.26)
2. Birthyear, and sex	0.30	(0.14, 0.46)	0.34	(0.05, 0.63)
3. Model 2 + birth state	0.18	(0.02, 0.33)	0.03	(-0.22, 0.27)
4. Model 3 + state condns	0.34	(0.11, 0.57)	-0.06	(-0.37, 0.26)
5. OLS estimates	0.09	(0.08, 0.10)	0.15	(0.14, 0.16)

# Evaluating Instruments

- Is the dependent variable independent of the instrument conditional on the endogenous variable?
- Over-identification tests, if you have multiple instruments
- Inequality constraints (for categorical endogenous variables)
- Evaluate the association between the instrument and the outcome across environments that modify the 1<sup>st</sup> stage association

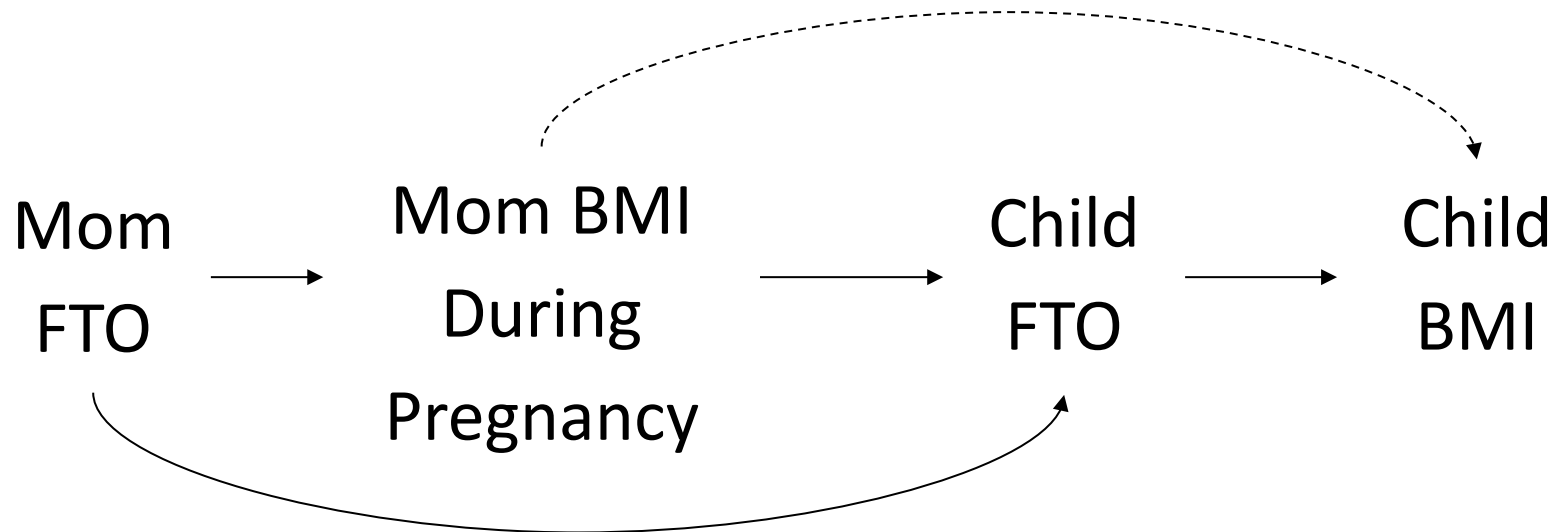


# Sensitivity Analyses

- Including education  $>13$  years
  - $\beta_{IV}$  (memory, model 3): 0.15 (-0.01, 0.31)
- Restricting to education  $> 13$  years
  - Instruments do not predict education or memory for individuals with  $>13$  years of school
  - $\beta_{IV}$  (memory, model 3): -1.04 (-3.70, 1.62)
- Inverse probability weighted for missing Memory (parental SES, self-report chronic condns at baseline)
  - $\beta_{IV}$  (memory, model 3): 0.19 (0.03, 0.36)

# Example 3: Maternal FTO as an IV for effect of mom's BMI on child's BMI

**Goal was to test developmental overnutrition hypothesis: exposure during gestation affects child BMI**

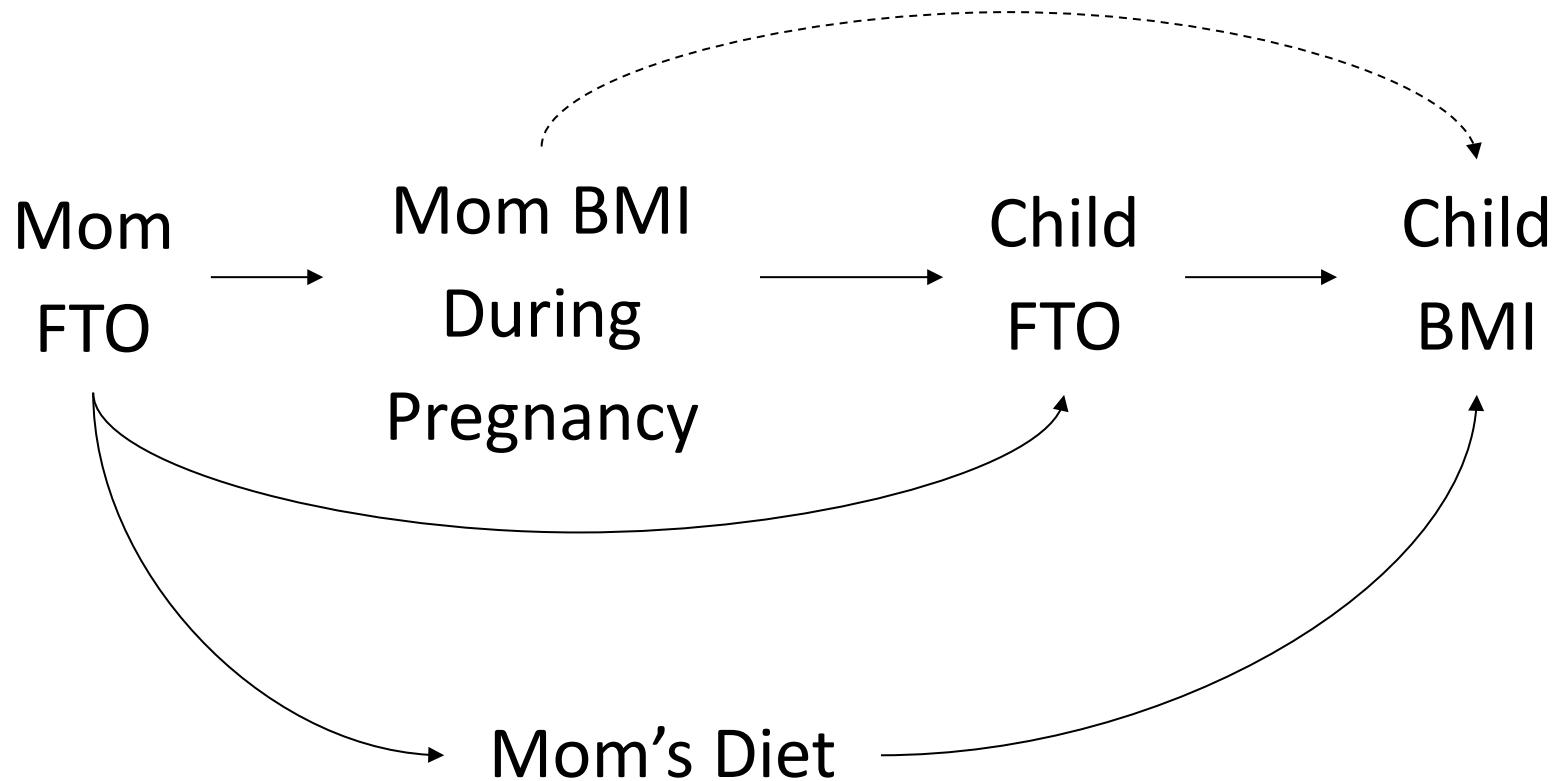


# IV effect estimates for Maternal BMI on Offspring total fat mass

	<b>OLS</b>	<b>IV</b>	<b>P-value for test of difference OLS vs IV</b>
Total Fat Mass	0.26 (0.23, 0.29)	-0.08 (-0.56, 0.41)	.17

From Lawlor PLoS Medicine 2008

# Example 3: Maternal FTO as an IV for effect of mom's BMI on child's BMI



# Doubling Instruments

- Do they have other pathways to the outcome?
  - *Quarter of birth*
- Is there a common cause of the instrument and the outcome?
  - *State of birth*
- Do they actually affect anyone's exposure?
  - *Tax policies*

# Thinking of Instruments, Creating Instruments

- Often ecological
- Policy changes
- Policy discontinuities
- Differences in “expert” opinion
- Encouragement designs: randomize the incentive
- **Ask: What is the process that determines exposure? Is any part of this process arbitrary/random?**
- Content matter experts are very valuable team members

# Conclusions

- Many important questions not convincingly answered with observational evidence
- Abandon the difficult questions? Or learn what we can from fraught methods?
- IV adds:
  - A way forward with observational data
  - Sometimes a parameter estimate of special interest
  - Pushes us to identify interventions that change exposures
- Not a replacement for evidence from observational research or RCTs, but a useful supplement

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